

MATHEMATICAL TREATMENT OF DUAL SCALING

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ABSTRACT

On the basis of the Preparations so far, we are now ready to look at a formal derivation of dual scaling. As you may have guessed from the many names for the quantification method and from the earlier discussion of the many ideas behind dual scaling, there are a number of ways to formulate it. This paper contains a detailed description of one of them. Some readers may find this paper too technical. Many formulas are presented in this paper occasionally without sufficient background information. However, once of the basic idea is understood, technical matters can be mastered through self-directed endeavor, and this paper should serve such guide by offering bird's eye view of mathematical treatment of dual scaling.

OPTIMAL SOLUTION

The Principle component analysis, in which the object is to determine weights for categories so as to maximize the variance of the composite scores, the analysis - of - variance approach, in which category weights are determined so as to maximize the between - row (column) sum of squares, relative to the total sum of squares, that is the squared correlation ratio, the bivariate correlation approach, in which the product - moment correlation between responses weighted by row weights and those by column weights maximized, and the simultaneous linear regression approach, in which row and column weights are determined so as to make the regression of rows on columns and the

regression of columns on rows simultaneously linear. This decomposition is a basis for the one - way analysis of variance and when data are continuous, the decomposition is uniquely determined in contrast when data are categorical, relative magnitudes of those sums of squares vary, depending on weights given to the categories. In other words, one can manipulate weights to change the relative magnitudes of Each sum of squares. Dual scaling chooses the weights in a particular way, the weights are determined in such a way that the ratio SS_b/SS_t that is the squared correlation ratio η^2 . be a maximum, In our task is to determine weights X_j so as to maximize η^2 , (1) In terms of matrix notation see equation (1), (2), (3) and (4)

	<i>Good</i>	<i>Averg</i>	<i>Poor</i>	
<i>White</i>	f_{11}	f_{12}	f_{13}	$y_1 = \sum x_j f_{1j}$
<i>Green</i>	f_{21}	f_{22}	f_{23}	$y_2 = \sum x_j f_{2j}$
<i>Brown</i>	f_{31}	f_{32}	f_{33}	$y_3 = \sum x_j f_{3j}$
Total	$\sum f_{j1}$	$\sum f_{j2}$	$\sum f_{j3}$	$\sum \sum x_j f_{ij} = y_t$

$$SS_b = \sum (y_j^2 / f_j) - y_t^2 / f_t$$

$$SS_w = \sum \sum y_{ij}^2 - \sum (y_j^2 / f_j)$$

$$SS_t = SS_b + SS_w$$

$$\frac{SS_b}{SS_t} = \eta^2$$

$$SS_t$$

If we have a matrix we have to equation (3,8)

F= (f_{ij}) :the Data matrix

f_r :the vector of row totals of F

f_c :the vector column totals of F

D_r :the diagonal matrix with row totals in the main diagonal

D_c : A diagonal diagonal matrix with column totals in the main diagonal¹

γ : a vector of weights for the rows of F

χ : a vector of weights for the columns of F

f_t : the sum of element of F, that is $\sum \sum f_{ij}$

Table 1

Teacher	Good	Average	Poor	Total
White	1	3	6	10
Green	3	5	2	10
Brown	6	3	0	9
	10	11	8	29

(2)

An illustrative example:

Let us follow the entire process of dual scaling, using the data in Table 1

That is the data used to illustrate the method of reciprocal averages. Let us assign unknown X_1 , X_2 and X_3 to three categories good, average and poor, respectively. Since the origin and the unit of quantified data are arbitrary, it is customary to see the sum of weighted responses and the sum of squares of weighted responses equal to zero and f_i

(the total number of responses), respectively:

$$\sum_{i=1}^n \sum_{j=1}^m f_{ij} X_j = 0 \quad \dots \quad (1)$$

$$\sum_{i=1}^n \sum_{j=1}^m f_{ij} X_j^2 = f_i \quad \dots \quad (2)$$

The choice of the origin by (1) makes the average of the weighted response zero, eliminating the correction term

The correction term

$$SS_t = 10X_1^2 + 11X_2^2 +$$

$$8X_3^2$$

(3)

$$SS_b = \frac{(x_1 + 3x_2 + 6x_3)^2}{10} + \frac{(3x_1 + 5x_2 + 2x_3)^2}{10} + \frac{(6x_1 + 3x_2)^2}{9} \quad \dots \quad (4)$$

2

(3),(4)

In the current example we have

$$F = \begin{pmatrix} 1 & 3 & 6 \\ 3 & 5 & 2 \\ 6 & 3 & 0 \end{pmatrix}$$

$$f_r = \begin{pmatrix} 10 \\ 10 \\ 9 \end{pmatrix}$$

,

$$f_c = \begin{pmatrix} 10 \\ 11 \\ 8 \end{pmatrix}$$

$$D_r = \begin{pmatrix} 10 & 0 & 0 \\ 0 & 10 & 0 \\ 0 & 0 & 9 \end{pmatrix}$$

,

$$D_c = \begin{pmatrix} 10 & 0 & 0 \\ 0 & 11 & 0 \\ 0 & 0 & 8 \end{pmatrix}$$

$$f_t = 29$$

$$y = \begin{pmatrix} y_1 \\ y_2 \\ y_3 \end{pmatrix}$$

,

$$X = \begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix}$$

$$1 = \begin{pmatrix} 1 \\ 1 \\ 1 \end{pmatrix}$$

$$I = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix}$$

³₍₈₎

$$f'_c x = 0 \quad \dots \quad (5)$$

$$x' D_C x = f_t \quad \dots \quad (6)$$

$$SS_t = x' D_C x \quad \dots \quad (7)$$

$$SS_b = x' F' D_r^{-1} F x \quad \dots \quad (8)$$

$${}^1 F' D_r^{-1} F x = (x_1, x_2, x_3) \begin{pmatrix} 1 & 3 & 6 \\ 3 & 5 & 3 \\ 6 & 2 & 0 \end{pmatrix} \begin{pmatrix} \frac{1}{10} & 0 & 0 \\ 0 & \frac{1}{10} & 0 \\ 0 & 0 & \frac{1}{9} \end{pmatrix} \begin{pmatrix} 1 & 3 & 6 \\ 3 & 5 & 2 \\ 6 & 3 & 0 \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \\ x_3 \end{pmatrix}$$

(5) Carry out the remainder of this numerical verification step by step . You will find the expression to be identical with Equation (4).

The object of dual scaling in this approach is to determine \mathcal{X} so as to maximize the squared correlation ratio η^2 , which is given by SS_b/SS_t .

This problem can be handled in two ways one is to maximize SS_b/SS_t in terms of \mathcal{X} , and the other is to maximize SS_b subject to the condition that $SS_t = f_t$ as indicated by equation (7) . Both lead to the same outcome. A widely used Procedure is the later one.

which is handled by the Well-Known Lagrange method of unknown multipliers .

Now , in the Lagrange method , the objective function for optimization is called Lagrange function which is in the present example given by :

$$Q(\mathcal{X}, \lambda) = SS_b - \lambda(SS_t - f_t) = x'F'D_r^{-1}Fx - \lambda(x'D_Cx - f_t)$$

(6)The function Q has two unknowns , \mathcal{X} and λ and we wish to determine \mathcal{X} and λ so as to maximize Q , Following the standard procedure

$$D_c^{-1}F'D_r^{-1}F = \begin{pmatrix} 0.5 & 0.38 & 0.12 \\ 0.35 & 0.40 & 0.25 \\ 0.15 & 0.33 & 0.50 \end{pmatrix}$$

To arrive at the standard form we will use the following transformation

$$B = D_r^{-\frac{1}{2}}FD_c^{-\frac{1}{2}} \quad \dots \quad (11)$$

$$B'B = D_c^{-\frac{1}{2}}F'D_r^{-1}FD_c^{-\frac{1}{2}}$$

And setting $D_c^{-\frac{1}{2}}x = w$

We obtain the following eigenequation in the standard form .

$$(B'B - \eta^2 I)w = 0 \quad \dots \quad (12)$$

Equation (12) can be written as

$$^5 (0.5000 - \eta^2)w_1 + 0.3623w_2 + 0.1342w_3 = 0$$

$$0.3623w_1 + (0.4000 - \eta^2)w_2 + 0.2985w_3 = 0$$

$$0.1342w_1 + 0.2985w_2 + (0.5000 - \eta^2)w_3 = 0$$

Eigenequation (12) has two trival equation

$$1: w = 0 \quad \dots \quad (13)$$

It is obvious that(17) is solution .But ,as we have already this solution is avoided by solving the characteristic equation that is

$$| B'B - \eta^2 I | = 0$$

for maximizing quadratic function, we set the partial derivatives of Q with respect to \mathcal{X} and λ equal to zero .

$$\frac{\partial Q}{\partial x} = 2F'D_r^{-1}Fx - 2\lambda D_Cx = 0 \quad \dots \quad (9)$$

$$\frac{\partial Q}{\partial \lambda} = x'D_Cx - f_t = 0$$

The first formula can be written as

$$F'D_r^{-1}Fx = \lambda D_Cx \quad \dots \quad (10)$$

If we premultiply both sides of this expression by x' we obtain

$$x'F'D_r^{-1}Fx = \lambda x'D_Cx$$

$$\lambda = \frac{x'F'D_r^{-1}Fx}{x'D_Cx} = \eta^2$$

(7)

Thus, the unknown multiplier λ is nothing but the squared correlation ratio that we want to maximize .

There for equation(9) can be written

$$(F'D_r^{-1}F - \eta^2 D_C)x = 0$$

In our example

Trival solution²

$$w = D_c^{-\frac{1}{2}}1 \quad \text{that is } \mathcal{X} = 1 \quad \dots \quad (14)$$

In this case the corresponding eigenvalue η^2 is 1 . This can easily be shown as fallows .first rewrite equation (12) as

$$B'Bw = \eta^2 w \quad \dots \quad (15)$$

Then if we substitute

$$w = D_c^{-\frac{1}{2}}1$$

For the left – hand side of equation (15) we obtain

$$B'Bw = D_c^{-\frac{1}{2}}F'D_r^{-1}FD_c^{-\frac{1}{2}}D_c^{-\frac{1}{2}}1$$

$$D_c^{-\frac{1}{2}}F'D_r^{-1}F1$$

$$F1 = f_r \text{ then } D_r^{-1}f_r = 1$$

$$D_c^{-\frac{1}{2}}F'1 = D_c^{-\frac{1}{2}}f_c$$

$$= D_c^{-\frac{1}{2}}1 \quad \dots \quad (16)$$

By substituting $w = D_c^{-\frac{1}{2}} \mathbf{1}$ for the right hand to equation (14) which gives $\eta^2 D_c^{-\frac{1}{2}} \mathbf{1}$

And compare this equation with equation (15) then $\eta^2 = 1$ when $w = D_c \mathbf{1}$, Where $\mathcal{X} = 1$

This allows us to put columns (categories) in (10) to even SS_w and $SS_b = 0$ then that $\mathcal{X} = 1$ amounts to putting all the columns (categories) in to one, making both SS_w and SS_b zeros, hence no information

to analyze for the move this "Solution" vector does not satisfy the condition that $F'_c \mathcal{X} = 0$.

Thus we must eliminate the second trivial solution from our computation. To carry out this task, the following theorem is relevant:

LaGrange's Theorem (Rao 1965): Let G be any square matrix of order n and rank $m > 0$, and p and q be column vector such that $P' G q$ is not zero.

Then, the residual matrix G_1 is exactly of rank $m-1$, where

$$G_1 = G - Gq (P' G q)^{-1} P' G \quad \dots \quad (17)$$

We can use this formula to eliminate the second trivial solution from $B' B$ of equation (11) is symmetric and that the trivial solution is an Eigen vector. Therefore, $p = q = a$, say, and $Ga = \lambda a$, an application of equation (17) to an Eigen equation yields the following formula for eliminating the effect of solution from an Eigen equation (5)

$$G_1 = G - \lambda a(a' a)^{-1} a' \quad \dots \quad (18)$$

If we normalize a to a^* by $a^* = \frac{a}{\sqrt{a' a}}$ that is $a^{*'} a^* = 1$ then

$$G_1 = G - \lambda a^* a^{*'} \quad \dots \quad (19)$$

This formula is used to extract the contribution of each successive solution in dual scaling and as such formula (19) is important to remember.

Applying formula (18) to our problem of eliminating the second trivial solution from the equation, we calculate the residual matrix,

indicated by C_1 , noting a in (18) is $D_c^{-\frac{1}{2}} \mathbf{1}$ as

$$C_1 = B' B - \frac{D_c^{-\frac{1}{2}} \mathbf{1} \mathbf{1}' D_c^{-\frac{1}{2}}}{f_t} \quad \dots \quad (20)$$

Note that in deriving formula (20) we also used the relations $\eta^2 = 1$ and

$W' W = 1 D_c^{-\frac{1}{2}} D_c^{-\frac{1}{2}} \mathbf{1} \mathbf{1}' = f_t$ The last term of the right-hand side of formula (20) in the current example is given as:

$$(8,9) \quad C_1 = \begin{pmatrix} 0.1552 & 0.0006 & -0.1742 \\ 0.0006 & 0.0207 & -0.0250 \\ -0.1742 & -0.0250 & 0.2241 \end{pmatrix}$$

Let us use $[1 \ 0 \ -1]$ as an initial vector b_0 then

$$C_1 b_0 = \begin{pmatrix} 0.3294 \\ 0.0256 \\ -0.3983 \end{pmatrix} = b_1$$

The largest absolute value in b_1 is 0.3983. Thus $\max(b_1) = k_1 = 0.3983$ and

$$b_1^* = \frac{b_1}{0.3983} = \begin{pmatrix} 0.8270 \\ 0.0643 \\ -1.000 \end{pmatrix} \quad \text{There fore}$$

Equation (6) the vector b_5^* must be rescaled

$$C_1 b_1^* = \begin{pmatrix} 0.3026 \\ 0.0268 \\ -0.3698 \end{pmatrix} = b_2, \quad K_2 = 0.3698$$

We repeat

$$C_1 b_1^* = b_{j+1}, \quad \frac{b_{j+1}}{|K_{j+1}|} = b_{j+1}^*$$

Until we reach

$$b_j^* = b_{j+1}^*$$

(10)

Since $b_4 = b_5$ the iterative process has converged to the solution and

$K_5 = \eta_1^2 = 0.3683$, That K_5 is the largest Eigen value to obtain the corresponding weight vector that satisfies

Equation (6) states that $X' D_c X = f_t$ But $X' D_c X = W' W$, and $f_t = 29$ There fore

$$W = \sqrt{\frac{29}{b_5^{*'} b_5^*}} \cdot b_5^* = \begin{pmatrix} 3.4029 \\ 0.3051 \\ -4.1626 \end{pmatrix}$$

The optimal weight vector \mathcal{X}_1 is given by:

$$(11,12) \quad \mathcal{X}_1 = D_c^{-\frac{1}{2}} w = \begin{pmatrix} \frac{-1}{\sqrt{10}} & 0 & 0 \\ 0 & \frac{1}{\sqrt{11}} & 0 \\ 0 & 0 & \frac{1}{\sqrt{8}} \end{pmatrix} \begin{pmatrix} 3.4029 \\ 0.3051 \\ -4.1626 \end{pmatrix} = \begin{pmatrix} 1.0761 \\ 0.0220 \\ -1.4717 \end{pmatrix}$$

Therefore, optimal Weights for categories good, average, and poor are 1.0761, 0.0920, and -104717.

We can start in same way as we have shown, except that we look at y for x and the between-column sum of squares for the between-row sum of squares .The squared correlation ratio η^2 is now defined as ratio of the between-column sum of squares to the total sum of squares

And this ratio is maximized with respect to y under the constraints $f' y=0$ and $y' D_r y = f_t$ Following the same procedure as before we obtain

$$(BB' - \eta^2 I) D_r^{-1} y = 0 \quad \dots (21)$$

Let η be the positive square root of η^2 . Then

$$y = \left(\frac{1}{\eta}\right) D_r^{-1} F X \quad \dots (22)$$

There are two trivial solution as before, $D_r^{-1} y = 0$ and $y=1$.The second trivial vector always yields $\eta^2 = 1$.

So every process is patrolled to the previous problem. Indeed, there exists complete symmetry or duality between the two problems .Thus, rather than repeating the same process let us present an important relation between x and y .

$$X = \left(\frac{1}{\eta}\right) D_c^{-1} F^{-1} y \quad \dots (23)$$

$$\begin{pmatrix} y_1 \\ y_2 \\ y_3 \end{pmatrix} = \frac{1}{\sqrt{0.3683}} \begin{pmatrix} \frac{1}{10} & 0 & 0 \\ 0 & \frac{1}{10} & 0 \\ 0 & 0 & \frac{1}{9} \end{pmatrix} \begin{pmatrix} 1 & 3 & 6 \\ 3 & 5 & 2 \\ 6 & 3 & 0 \end{pmatrix} \begin{pmatrix} 1.071 \\ 0.0920 \\ -104717 \end{pmatrix}$$

$$= \begin{pmatrix} -1.2322 \\ 0.1227 \\ 1.2326 \end{pmatrix}$$

Conclusion:

We can notice that $X_1 < X_2 < X_3$ and $y_1 < y_2 < y_3$ i.e. we find that the best teach is the third with the best optimal weight for categoriese.t

From the outset of dual scaling research the mean orientation has been to extend its applicability to a wider variety of categorical data .As arousal data (i.e. the main data types for correspondence analysis) , but also purid comparison rank order , successive categories , sorting and multiway data matrices .

In addition to this wider applicability an analogue of discriminant analysis for categorical data has been developed under the name forced classification and the procedure of generalized forced classification is also available. Other computational procedures were also developed Including the method of successive data modification (S D M) for quantification of ordered categories , partially optimal scaling for

data with preassigned weight or mixed (categorical plus continuous) data, and the piecewise method of reciprocal average for handling a large number of multiple-choice items .The introduction and discussion of many of these procedures would certainly be enough for one book- another reason for staying within the familiar territory of dual scaling .

Consequently there may be an excessive number of references to those studies conducted in Toronto.

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